



***PARAMETRIC SURVIVAL MODELS WITH TIME-DEPENDENT
COVARIATE FOR MIXED CASE INTERVAL-CENSORED DATA***

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TIME-DEPENDENT COVARIATE FOR MIXED CASE
INTERVAL-CENSORED DATA**

By

KAVEH KIANI

Thesis Submitted to the School of Graduate Studies, Universiti Putra
Malaysia, in Fulfilment of the Requirements for the Degree of Doctor
of Philosophy

May 2012



DEDICATIONS

To

My parents

Who have supported me all the way since the beginning of my life

My parents-in-law

For their great encouragement and support

My wife, Marjan

Who has been a great source of love and motivation

My daughter, Niki

For her great patience

Abstract of thesis presented to the Senate of Universiti Putra Malaysia in
fulfilment of the requirement for the degree of Doctor of Philosophy

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May 2012

Chair: Jayanthi Arasan, PhD

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The aim of this research is to analyze parametric survival models in the presence of left, right, interval and doubly interval censored data with time-dependent covariates. In this research we utilize and extend two important parametric survival models, the Gompertz and the exponential, to accommodate these censoring schemes and time-dependent covariates.

The analysis starts with the extension of the Gompertz model to incorporate time-dependent covariates in the presence of right-censored data. Then, the performance of the model is compared with the fixed covariate model. Following that, comparison is made when a fixed covariate model was fitted wrongly to a data set with time-dependent covariate. In addition, two methods of constructing confidence intervals, the Wald and jackknife are explored for the parameters of this model. Conclusions are drawn based on the coverage probability study.

In the next step, the Gompertz model is further extended to incorporate time-

dependent covariates with left, right and interval censored data as well as uncensored data. The model is then investigated thoroughly at dependent and independent covariate levels through a comprehensive simulation study. Following that, the model is compared with a fixed covariate model. Then, two methods of constructing confidence intervals the Wald and likelihood ratio are investigated for the parameters of the model and conclusions are drawn based on the coverage probability study.

Finally, a parametric survival model that accommodates doubly interval-censored data with time-dependent covariates is developed and studied. In order to formulate this censoring scheme let V and W be the times of two related consecutive events where both of them are interval-censored and $V \leq W$. Then, the survival time of interest could be defined as, $T = W - V$. Here it is assumed that the time to the first event, V , and the survival time, T , follow the exponential distribution (special case for Gompertz distribution).

In order to get to this final model, we had to explore three separate models in advance. Firstly, a simple model consisting doubly interval-censored data without any covariate was studied. Following that, a model with doubly interval-censored data and fixed covariates is considered. Lastly, a model with fixed covariates is studied where some of the covariates affect T and the others affect V . All these models are studied by the simulation study and two methods of constructing confidence intervals, the Wald and jackknife are explored for the parameters of the models.

The results indicate that the Gompertz model with left, right and interval censored data with a time-dependent covariate works rather well despite its complexity. Similarly, although doubly interval-censored data with a time-dependent covariate requires more computational effort, the model will perform well if both V and T are exponentially distributed.

Abstrak tesis yang dikemukakan kepada Senat Universiti Putran Malaysia
sebagai memenuhi keperluan untuk ijazah Doktor Falsafah

**MODEL MANDIRIAN BERPARAMETER DENGAN KOVARIAT
YANG BERSANDAR KEPADA MASA UNTUK KES DATA
TERTAPIS SELANG BERCAMPUR**

Oleh

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May 2012

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Matlamat kajian ini adalah untuk menganalisis model mandirian berparameter dengan data tertapis kiri, kanan, selang dan selang berganda dengan kovariat yang bersandar kepada masa. Dalam kajian ini, kita mengguna pakai dan melanjutkan dua model mandirian berparameter yang penting, *Gompertz* dan eksponen, untuk menampung skim penapisan tersebut dan kovariat yang bersandar kepada masa. Kajian bermula dengan lanjutan model *Gompertz* untuk mengambil kira kovariat-kovariat yang bersandar kepada masa dengan data tertapis kanan. Kemudian, prestasi model ini dibandingkan dengan dengan model berkovariat tetap. Seterusnya perbandingan dilakukan apabila model berkovariat tetap yang digunakan dan bukannya model sebenar dengan kovariat yang bersandar kepada masa. Di samping itu, dua kaedah membina selang keyakinan, iaitu *Wald* dan *jackknife* telah dikaji bagi parameter-parameter model ini. Kesimpulan telah dibuat berdasarkan kajian kebarangkalian liputan.

Dalam langkah seterusnya, model *Gompertz* telah dilanjutkan lagi untuk mengambil kira kovariat-kovariat yang bersandar kepada masa dengan data tertapis kiri,

kanan dan selang disamping data tidak tertapis. Model ini kemudiannya telah dikaji secara menyeluruh pada tahap kovariat yang bersandar dan tidak bersandar melalui kajian simulasi yang komperhensif. Berikutan itu, model ini dibandingkan dengan model berkovariat tetap. Kemudian, dua kaedah membina selang keyakinan iaitu *Wald* dan kemungkinan nisbah dikaji bagi parameter-parameter model ini dan kesimpulan telah dibuat berdasarkan kajian kebarangkalian liputan.

Akhirnya, model mandirian berparameter yang mengambilkira data tertapis selang berganda dengan kovariat-kovariat yang bersandar kepada masa telah dikira dan dikaji. Bagi tujuan untuk memformulasikan skim data tertapis ini, biar V dan W menjadi masa-masa bagi kejadian berturutan yang berkaitan dimana keduanya adalah data tertapis selang dan $V \leq W$. Kemudian, masa mandirian yang dikehendaki boleh di definisikan sebagai, $T = W - V$. Disini, adalah dianggap bahawa masa ke kejadian yang pertama, V dan masa mandirian, T , bertaburan eksponen (kes khas bagi *Gompertz*).

Untuk mencapai model terakhir, kita terpaksa mengkaji tiga model yang berasin-gan dahulu. Pertama sekali, model ringkas yang terdiri daripada data tertapis selang berganda tanpa sebarang kovariat telah dikaji. Seterusnya, model dengan data tertapis selang berganda dan satu kovariat tetap telah diambilkira. Akhir sekali, model dengan kovariat-kovariat tetap telah dikaji dimana sebahagian daripada kovariat tersebut memberi kesan kepada V . Kesemua model-model ini telah dikaji dengan menggunakan kajian simulasi dan dua kaedah pembinaan selang keyakinan, *Wald* dan *jackknife* telah dikaji bagi parameter-parameter model ini. Keputusan menunjukkan bahawa model *Gompertz* dengan data tertapis kiri, kanan dan selang dengan kovariat yang bersandar kepada masa berfungsi dengan agak baik walaupun ianya kompleks. Begitu juga, walaupun data tertapis selang yang bersandar kepada masa memerlukan usaha komputasi yang lebih, model tersebut berfungsi dengan baik apabila kedua dua V dan T bertaburan eksponen .

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I certify that a Thesis Examination Committee has met on **May 2012** to conduct the final examination of **Kaveh Kiani** on his thesis entitled "**Parametric Survival Models with Time-Dependent Covariate for Mixed Case Interval-Censored Data**" in accordance with the Universities and University Colleges Act 1971 and the Constitution of the Universiti Putra Malaysia [P.U.(A) 106] 15 March 1998. The Committee recommends that the student be awarded the **Doctor of Philosophy**.

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DECLARATION

I declare that the thesis is my original work except for quotations and citations which have been duly acknowledged. I also declare that it has not been previously, and is not concurrently, submitted for any other degree at Universiti Putra Malaysia or at any other institution.



KAVEH KIANI

Date:

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LIST OF ABBREVIATIONS

AC	Anti-Conservative
Asy	Asymmetrical
C	Conservative
CI	Confidence Interval
CP	Censoring Proportion
DIC	Doubly Interval-Censored
EDIC	Exponential Model with Doubly Interval-Censored Data
EDICF1	Exponential Model with Doubly Interval-Censored Data and Fixed Covariates, Case 1
EDICF2	Exponential Model with Doubly Interval-Censored Data and Fixed Covariates, Case 2
EDICTD	Exponential Model with Doubly Interval-Censored Data and Time-Dependent Covariates
GICF	Gompertz Model with Interval-Censored Data and Fixed Covariates
GICTD	Gompertz Model with Interval-Censored Data and Time-Dependent Covariates
GRCF	Gompertz Model with Right-Censored Data and Fixed Covariates
GRCTD	Gompertz Model with Right-Censored Data and Time-Dependent Covariates
IC	Interval-Censored
LC	Left-Censored
MLE	Maximum Likelihood Estimate
OE	Observed Exactly
RC	Right-Censored
RMSE	Root Mean Square Error
SE	Standard Error
TD	Time-Dependent

CHAPTER 1

INTRODUCTION

1.1 Background Study

In this section, survival time data analysis along with some basic definitions, models and concepts that will be used in subsequent chapters are introduced. These topics will not be discussed in detail.

1.1.1 Survival Time Data Analysis

Survival time or in some literature, life time, time-to-event and failure time, T , is the time from an initiating event (birth, marriage, affecting with a virus, enroll for a PhD. program) for a subject to some final event (death, divorce, onset of the diseases, graduation). T is the non-negative random variable and usually continuous, unless stated otherwise.

Survival time data analysis, briefly, survival analysis is a collection of statistical methods for analyzing survival data where, the focus of these methods is to describe the distribution of T on a population and relationship between T and some covariates. A serious analytical problem in analyzing survival time data arises when a portion or even all $t_i, i = 1, 2, \dots, n$, are censored. The i^{th} subject's survival time, t_i , is censored when we do not know its exact value due to one of these four main reasons, the subject is lost to follow-up, the subject withdraws from the study, the subject does not experience the event before the study ends and the subject is not under continuous observation. However, we may have some partial information regarding the subject's survival time.

In survival analysis it is essential to have clear and explicit definitions of the time origin and endpoint. Following Kalbfleisch and Prentice (2002), time origin is

the time from which survival is measured, $t = 0$. In some instances, time origin and endpoint could be the birth and death of the subject. Another example is infection with a virus and onset of the disease. In clinical trials time origin could also be assumed to be the time when all of the subjects in the trial had been event free.

The statistical analysis and modeling of survival time data are usually done by applying various kinds of non-parametric, semi-parametric or parametric models. Non-parametric models do not assume any parametric assumption about the form of the survivor function, $S(t)$.

On the other hand, a semi-parametric model is partially attached to some parametric assumption for $S(t)$. As Sasieni (2005) explains, "There is no widely accepted rigorous definition of a semi-parametric model. Informally we will call a model semi-parametric if it is not fully parametric but has a finite dimensional parameter of interest". In contrast, parametric models completely assume a fully parametric form for $S(t)$. More discussions about these three models are given in Sections 1.1.4 to 1.1.6.

1.1.2 Different Censoring Mechanisms

The i^{th} subject is observed exactly (OE) at t_i if the event of interest occurs at this time and this information is available to the observer. This subject is left-censored (LC) when $t_i \in (0, l_i]$ or, the subject has met the event at unknown time prior to l_i and after time origin and it is right-censored (RC) when $t_i \in (r_i, \infty)$ or, the subject has been event free at the last known time r_i . Finally, the i^{th} subject is interval-censored (IC) if instead of observing t_i only an interval $(t_{L_i}, t_{R_i}]$ is observed where $t_i \in (t_{L_i}, t_{R_i}]$ and $t_{L_i} \leq t_{R_i}$.

Following Lawless (2003), there are two types of RC data, type-I and type-II. To describe type-I right censoring scheme let us assume that each subject has

a potential censoring time c_i and an event time t_i . The i^{th} subject is RC if $c_i < t_i$; otherwise, subject is OE. This scheme usually occurs when study period is determined in advance. Type-II right censoring scheme arises when the study is terminated after first k failures where $k \leq n$ and n is number of subjects. Type-I RC data is referred to as RC data in this study.

IC data occur when subjects are under discontinuous observations/inspections. In this case, T is not always OE or RC. IC data are more common in fields such as, sociology, economic, biology and epidemiology. Occurrence of IC data in medical studies is also very common when subjects are under scheduled follow-ups according to a predetermined calendar time for example weekly, monthly or yearly. In this study the term "actual inspection time (AC)" is used to denote the time when the subject was inspected.

Following Sun (2006) and Schick and Yu (2000) the six types of IC data are, case-I, case-II, case-k, mixed case and doubly IC data and panel count data.

Case-I IC data or current status data is referred to the IC data when all the $(T_L, T_R]$ intervals include either 0 or ∞ (Groeneboom and Wellner, 1992; Huang, 1996 and Schick and Yu, 2000). In this scheme, the i^{th} subject is inspected only once at ac_{i1} where ac_{i1} will be left or right censoring time. Case-I IC data differ from RC data and LC data because there is not any OE data in this case. This type of IC data could be represented by

$$\left\{ ac_{i1}, \delta_{i1} = I(t_i \leq ac_{i1}), \delta_{i2} = I(t_i > ac_{i1}), i = 1, \dots, n \right\},$$

where I is an indicator variable and ac_{i1} is the actual inspection time for the i^{th} subject. In this case

$$(t_{L_i}, t_{R_i}] = \begin{cases} (0, ac_{i1}], & \delta_{i1} = 1, \\ (ac_{i1}, \infty), & \delta_{i2} = 1. \end{cases} \quad (1.1)$$

In case-II IC data each subject is inspected twice. As a result, there is at least one finite interval $(T_L, T_R]$ belonging to $(0, \infty)$ (Groeneboom and Wellner, 1992; Huang and Wellner, 1997; Sun, 2005; Schick and Yu, 2000). This type of IC data could be represented by

$$\left\{ ac_{i1}, ac_{i2}, \delta_{i1} = I(t_i \leq ac_{i1}), \delta_{i2} = I(ac_{i1} < t_i \leq ac_{i2}), \delta_{i3} = I(t_i > ac_{i2}), i = 1, \dots, n \right\},$$

where ac_{i1} and ac_{i2} are two actual inspection times and $ac_{i1} \leq ac_{i2}$. In this case

$$(t_{L_i}, t_{R_i}] = \begin{cases} (0, ac_{i1}], & \delta_{i1} = 1, \\ (ac_{i1}, ac_{i2}], & \delta_{i2} = 1, \\ (ac_{i2}, \infty), & \delta_{i3} = 1. \end{cases} \quad (1.2)$$

Case-k IC data occurs when there are k actual inspection times for subjects, $ac_{i1} \leq ac_{i2} \leq \dots \leq ac_{ik}$ where, k is a fixed number (Schick and Yu, 2000; Wellner, 1995). This type of IC data could be represented by

$$\left\{ ac_{ij}, k, \delta_{ij} = I(ac_{i(j-1)} < t_i \leq ac_{ij}), \delta_{i(k+1)} = I(t_i > ac_{ik}), i = 1, \dots, n, j = 1, \dots, k \right\},$$

where $ac_{i0} = 0$. In this case

$$(t_{L_i}, t_{R_i}] = \begin{cases} (0, ac_{i1}], & \delta_{i1} = 1, \\ (ac_{i(j-1)}, ac_{ij}], & \delta_{ij} = 1 \text{ and } 1 < j \leq k, \\ (ac_{ik}, \infty), & \delta_{i(k+1)} = 1. \end{cases} \quad (1.3)$$

Mixed case IC data occurs when there is k_i actual inspection times for the i^{th} subject or k is a random number, $ac_{i1} \leq ac_{i2} \leq \dots \leq ac_{ik_i}$ (Schick and Yu, 2000;

Wellner, 1995). This type of IC data could be represented by

$$\left\{ ac_{ij}, k_i, \delta_{ij} = I(ac_{i(j-1)} < t_i \leq ac_{ij}), \delta_{i(k_i+1)} = I(t_i > ac_{ik_i}), j = 1, \dots, k_i, i = 1, \dots, n \right\},$$

where $ac_{i0} = 0$. In this case

$$(t_{L_i}, t_{R_i}] = \begin{cases} (0, ac_{i1}], & \delta_{i1} = 1, \\ (ac_{i(j-1)}, ac_{ij}], & \delta_{ij} = 1 \text{ and } 1 < j \leq k_i, \\ (ac_{ik_i}, \infty), & \delta_{i(k_i+1)} = 1. \end{cases} \quad (1.4)$$

Mixed case IC data is very common in medical trials because the number of actual inspection times varies from patient to patient. It is clear that representation (1.4) is the general form of the representations (1.1), (1.2) and (1.3). In this study the term "IC" data is used to refer to "mixed case IC" data. More details about IC data is presented in Chapters 2 and 4.

Doubly interval-censored (DIC) data is the fifth type of IC data. DIC data often arises in the follow-up studies where T is the elapsed time between two related events where both events are IC (Gruttola and Lagakos, 1989; Sun, 2004). If V is the time to the first event and W is the time to the second event then

$$V \in (V_L, V_R], W \in (W_L, W_R] \text{ and } T = W - V,$$

where $V_L \leq V_R$ and $W_L \leq W_R$. DIC data include usual RC and IC survival time data as special cases. For example DIC data reduce to IC data if V is OE and W is IC and DIC data reduce to RC data if V is OE and W is RC. DIC data is discussed more in Chapters 2 and 5.

Finally, panel count data is another form of IC data where in the presence of interval censoring, event of interest occurs more than one time during follow-ups and researcher is interested to know the exact number of these occurrences

(Lawless and Zhan, 1998; Thall, 1988). This type of IC data is not explored in this study. For more discussions see Sun (2006).

1.1.3 Independent Censoring

Independent censoring is a realistic assumption in survival analysis. Independent or random censoring is a censoring mechanism that is independent of the subjects' survival time distribution. For instance, consider two subjects that are under trial at time t where one of them has a higher failure risk. Under independent censoring assumption, both subjects have a same probability to be RC at this time.

An independent censoring mechanism is also equivalent to the non-informative censoring. However, the reverse may not always true (Sun, 2006; Betensky, 2000; Oller et al., 2004). In contrast, dependent censoring and informative censoring are always the same.

For RC data assume $S(t)$, $f(t)$, $G(c)$ and $g(c)$ represent survivor function and probability density function of survival time T and right censoring time C . If censoring process produces a potential censoring time c_i for the i^{th} subject and t_i is the survival time of this subject then observed survival time would be $ot_i = \min(t_i, c_i)$. If $ot_i = c_i$ the indicator variable is $\delta_{R_i} = 1$ and $\delta_{R_i} = 0$ if $ot_i = t_i$. Dependent (informative) censoring for RC data scheme could be represented by

$$\begin{aligned} Pr(ot_i = t, \delta_{R_i} = 1) &= Pr(c_i = t, t_i > t) = Pr(c_i = t)Pr(t_i > t|c_i) = g(t)S(t|c_i), \\ Pr(ot_i = t, \delta_{R_i} = 0) &= Pr(t_i = t, c_i \geq t) = Pr(c_i \geq t|t_i)Pr(t_i = t) = G(t|t_i)f(t), \end{aligned}$$

and for independent (non-informative) censoring is

$$\begin{aligned} Pr(ot_i = t, \delta_{R_i} = 1) &= Pr(c_i = t, t_i > t) = Pr(c_i = t)Pr(t_i \geq t) = g(t)S(t), \\ Pr(ot_i = t, \delta_{R_i} = 0) &= Pr(t_i = t, c_i \geq t) = Pr(c_i > t)Pr(t_i = t) = G(t)f(t). \end{aligned}$$

For IC data assume $S(t)$ and $f(t)$ represent survivor function and probability density function of survival time T and $G(t_L, t_R)$ represents joint survivor function of left and right censoring time. If censoring process produces t_{L_i} and t_{R_i} as a left and right censoring times and t_i be the survival time of the subjects then dependent (informative) censoring for IC data scheme could be represent by

$$Pr(t_{L_i} = l, t_{R_i} = r | t_i = t) = Pr(l < t_i \leq r | t_{L_i} = l, t_{R_i} = r)G(l, r),$$

and for independent (non-informative) censoring is

$$Pr(t_{L_i} = l, t_{R_i} = r | t_i = t) = Pr(l < t_i \leq r).$$

For DIC data independent censoring assumption means that survival time, T , and time to the first event, V , are independent. Independent censoring for RC, IC and DIC data schemes are assumed throughout the thesis.

1.1.4 Parametric Survival Models

In parametric survival models T usually follows a continuous distribution function, $f(t)$ where the cumulative distribution function is $F(t) = \Pr(T \leq t)$, survivor function is $S(t) = 1 - F(t) = \Pr(T > t)$, hazard function is $h(t) = f(t)/S(t)$ and cumulative hazard function is $H(t) = \int_0^t h(v)dv$. When one of these five functions is known other four functions could be specified. As a result, a parametric survival model is a model with specified $S_{\theta}(t)$ function where, θ is the vector of parameters.

Some widely used distributions to model T are the exponential, Weibull, Gompertz, log-normal, log-logistic, generalized gamma and generalized F . The existences of different censoring mechanisms and different type of covariates have motivated researchers to introduce new parametric models and methods or extend existing parametric models to accommodate these components. Some of the most

important advantages of the parametric models are existence of straightforward methods to obtain the maximum likelihood estimation (MLE) of the parameters, confidence intervals (CI) and hypothesis testing procedures.

The exponential and Gompertz are two models which are explored in this research. The exponential model has a constant hazard rate

$$h(t) = \lambda,$$

where $\lambda > 0$ is a parameter. The survivor function of the model is

$$S(t) = \exp(-\lambda t),$$

and the probability density function is

$$f(t) = \lambda \exp(-\lambda t).$$

The hazard rate of this distribution, λ , is constant or independent of time. Constant hazard rate or lack of aging or memory less property of the exponential distribution implies that

$$Pr(T \geq k) = Pr(T \geq k + t | T \geq k).$$

This implies that the chance for a new born subject to survive at least up to time k is equal to chance of a k age subject to survive additional t units. Because of the memory less property this distribution has found limited use in studies where subject's aging process play an important role in its survival. This distribution is frequently used in the simulation studies in order to explore new and complicated models.

The Gompertz model was introduced by Gompertz (1825) as a model for human

mortality. Recently, it has found more application in fields such as biology and demography. The hazard function of the Gompertz model is

$$h(t) = \lambda \exp(\gamma t),$$

The scale parameter is $\lambda > 0$ and the shape parameter is γ . The survivor function of the model is

$$S(t) = \exp \left[\frac{\lambda}{\gamma} (1 - e^{\gamma t}) \right]$$

and the probability density function is

$$f(t) = \lambda \exp(\gamma t) \times \exp \left[\frac{\lambda}{\gamma} (1 - e^{\gamma t}) \right].$$

The properties of the Gompertz distribution is presented in Johnson et al. (1994) and recently many authors have done studies on different characteristics and statistical methodology of Gompertz distribution, for instance, Makany (1991) and Chen (1997). The exponential model is the special case of the Gompertz model. In other words, Gompertz model could be reduced to the exponential model when the shape parameter γ is assumed to be zero.

1.1.5 Non-Parametric Survival Models

Non-parametric methods have been discussed the most in the analysis of IC data. Kaplan and Meier (1958) derived an estimation for the $S(t)$ in the presence of RC data which is often referred to as the Kaplan-Meier or product-limit estimate. Thus, a non-parametric estimate for the cumulative hazard function has been proposed by Nelson (1969, 1972). Ayer et al. (1955) and van Eeden (1956) were the first to introduce non-parametric maximum likelihood estimates (MLE) of a distribution function based on case-I IC data. Non-parametric likelihood functions with other types of IC data are more analytically and practically com-

plicated than that RC and case-I IC data and MLE of a $S(t)$ function does not have a closed form and should be calculate via an iterative algorithm.

Peto (1973) and Turnbull (1976) studied the estimations based on the case-II IC data. Suppose that $t_i \in (t_{Li}, t_{Ri}]$ where, $i = 1, 2, \dots, n$ and t_i 's are iid with survivor function $S(t)$. Then the likelihood function is

$$L = \prod_{i=1}^n \left\{ S(t_{Li}) - S(t_{Ri}) \right\}. \quad (1.5)$$

If, $0 < s_0 < s_1 < s_2 < \dots < s_{m-1} < s_m = \infty$ denote the ordered elements of the set $\left\{ 0, t_{Li}, t_{Ri}, \infty, i = 1, 2, \dots, n \right\}$ then define an indicator variable $\alpha_{ij} = I(S_j \in (t_{Li}, t_{Ri}])$ and $p_j = S(s_{j-1}) - S(s_j)$, $j = 1, 2, \dots, m$. Expression (1.5) can be rewritten as

$$L(P) = \prod_{i=1}^n \sum_{j=1}^m \alpha_{ij} p_j.$$

MLE of this function could be determine by applying one of the iterative procedures like, self-consistency algorithm of the EM algorithm by Dempster (1977), ICM algorithm by Groeneboom and Wellner (1992) and later modified by Jongbloed (1998) and EM-ICM algorithm proposed by Wellner and Zhan (1997).

Wellner (1995), Groeneboom (1996) and Huang and Wellner (1997) studied the non-parametric MLE of the case-k IC data and Schick and Yu (2000) and Van der Vaart and Wellner (2000) discussed non-parametric MLE of the mixed case IC data.

Non-parametric approaches to estimate $S(t)$ in the presence of DIC data usually use the self-consistency idea, see Sun (2006). There are three estimation procedures at the base of this idea. Firstly, a method based on the ML approach has been proposed by Gruttola and Lagakos (1989). Secondly, a two-step procedure as a simplification of the first method has been proposed by Gómez and Lagakos (1994) and finally, a conditional likelihood-based approach proposed by

Sun (1997). All these three methods can also be regarded as a generalization of the self-consistency algorithm.

1.1.6 Semi-Parametric Survival Models

Two major semi-parametric models are Cox proportional hazard model (PH) and accelerated failure time (AFT) or log-rank model. Hazard function of the PH model is

$$h(t, \mathbf{X}) = h_0(t) \exp(\boldsymbol{\beta}'\mathbf{X}),$$

and for the AFT model is

$$h(t, \mathbf{X}) = h_0(te^{-\boldsymbol{\beta}'\mathbf{X}}) \exp(-\boldsymbol{\beta}'\mathbf{X}),$$

where \mathbf{X} is the vector of the p covariates, $\mathbf{X}=(x_1, x_2, \dots, x_p)$, and $\boldsymbol{\beta}$ is the vector of p parameters, $\boldsymbol{\beta}=(\beta_1, \beta_2, \dots, \beta_p)$ and $h_0(t)$ is an unspecified baseline hazard function.

Three types of methods have discussed in literature for analyzing IC data by the PH model. Firstly, use an imputation method to reduce IC data to RC data and then use specified methods for the RC data for example see, Satten (1996), Goggins et al. (1998) and Pan (2000a). Secondly, estimating the unspecified baseline hazard function, $h_0(t)$, non-parametrically and then estimate covariate's parameters, Finkelstein (1986) and Goetghebeur and Ryan (2000). Thirdly and finally, using regression splines or local likelihood smoothing to estimate baseline hazard function, $h_0(t)$, and then estimate rest of the parameters where, Kooperberg and Clarkson (1997) and Betensky and Finkelstein (1999) are pioneers of these two methods.

Analyzing IC data with AFT model is less developed compared to the PH model and some of the main works were given by Rabinowitz et al. (1995), Betensky et

al.(2001) and Komárek et al. (2005).

1.1.7 Time-Dependent Covariate

Many parametric regression models have been developed to analyze relationship between T as an outcome variable and covariates in the presence of censored data. Covariates can be either fixed or time-dependent (TD). Fixed covariates are measured at the start of study and stay constant over the study period, for instance, gender or race. TD covariates on the other hand vary over time, for example, age of subjects, blood pressure and cholesterol level. Cox (1972) has introduced idea of TD covariates and he has suggested the use of TD covariate in the PH regression models. Idea behind using the TD covariates implies that the history of a TD covariate process up to time t should be incorporated into the model to assess the full effect of the covariate on the T , because this history may influence the rate of survival over time.

Following Kalbfleisch and Prentice (2002), TD covariates are categorized as either external (exogenous) or internal (endogenous). External TD covariate may influence the observed survival time of a subject, t , but the covariate's path after t and t are independent. For instance, the effect of air pollution on subjects heart attack. In contrast, the path of the internal TD covariate is dependent on the survival status of a subject or the TD covariate process is generated by the subject for example, blood pressure of a subject. This covariate is observable as long as the subject is under observation.

1.2 Objectives

The aim of this research is to obtain parametric survival models for survival time data with LC, RC, IC, DIC and OE data in presence of TD covariates and obtain various CI estimates for the parameters of these models. In this research we will

utilize and extend two important parametric survival models, the Gompertz and the exponential to accommodate these censoring mechanisms and TD covariates. As mentioned before the exponential model is the special case of the Gompertz model. The exponential model is used in Chapter 5 to establish a base for future extensions with other parametric models include Gompertz.

Three main models will be explored,

- Gompertz model with RC data and TD covariates (GRCTD),
- Gompertz model with IC data and TD covariates (GICTD),
- Exponential model with DIC data and TD covariates (EDICTD).

In order to achieve these final models other models must be explored in advance.

The models are

- Gompertz model with RC data and fixed covariates (GRCF),
- Gompertz model with IC data and fixed covariates (GICF),
- Exponential model with DIC data (EDIC),
- Exponential model with DIC data and fixed covariates, Case 1 (EDICF1),
- Exponential model with DIC data and fixed covariates, Case 2 (EDICF2).

The main objectives of this research are as follows:

- To extend the Gompertz model to incorporate TD covariates in the presence of RC data or obtaining GRCTD model and conduct simulation to study the bias, standard error (SE) and root mean square error (RMSE) of the parameter estimates.

- To extend the Gompertz model to incorporate TD covariates in the presence of IC data or obtaining GICTD model and conduct simulation to study the bias, SE and RMSE of the parameter estimates.
- To extend the exponential model to incorporate TD covariates in the presence of DIC data or obtaining EDICTD model and conduct simulation to study the bias, SE and RMSE of the parameter estimates.
- Obtain GRCF, GICF, EDIC, EDICF1 and EDICF2 models and study performance of the models.
- Evaluate the performance of the the Wald, jackknife and likelihood ratio CI estimate methods for these models and draw the results based on the coverage probability study.

1.3 Outline of Thesis

This thesis is organized into six chapters. Chapter 2 provides a review of related literature to the current work. Special consideration in this literature review is the research conducted on parametric models with RC, IC and DIC data and parametric models with TD covariate.

Chapter 3 starts with the extension of the Gompertz model to incorporate TD covariates in the presence of RC data (GRCTD model). Then, the performance of the model is compared with the fixed covariate model (GRCF model). Following that, comparisons are made when a fixed covariate model is used instead of the true TD covariate model. In addition, two methods of constructing CIs, the Wald and jackknife are explored for the parameters of the model. Conclusions are drawn based on the coverage probability study.

Chapter 4 concentrates on the extension of the Gompertz model to incorporate TD covariates with LC, RC, IC data as well as OE data (GICTD model). The

model is then investigated thoroughly at dependent and independent covariate levels through a comprehensive simulation study. Following that, the model is compared with a fixed covariate model (GICF model). Then, two methods of constructing CIs the Wald and likelihood ratio are investigated for the parameters of the model and conclusions are drawn based on the coverage probability study.

Chapter 5 investigates a parametric survival model that accommodates DIC data with TD covariates (EDICTD model). In order to achieve this final model three models are explored in advance. Firstly, a simple model consisting DIC data without any covariate is studied (EDIC model). Following that, a model with DIC data and fixed covariates is considered where all the covariates affect T (EDICF1 model). Lastly, a model with fixed covariates is studied where some of the covariates affect T and the others affect V (EDICF2 model). All these models are studied by the simulation study and two methods of constructing confidence intervals, the Wald and jackknife are explored for the parameters of these models.

Finally, Chapter 6 summarizes the study and offers some recommendations for future research.

It should be mentioned all simulation studies were carried out by using the FORTRAN[®], (FTN95), programming language.

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LIST OF PUBLICATIONS

The following papers are extracted from the current thesis:

1. Kiani, K., Arasan, J. (2012). Gompertz model with time-dependent covariate in the presence of interval-, right- and left-censored data. *Journal of Statistical Computation and Simulation*, 1-19, doi: 10.1080/00949655.2012.662979, URL: <http://dx.doi.org/10.1080/00949655.2012.662979>.
2. Kiani, K., Arasan, J. and Midi, H. (2012). Interval Estimations for Parameters of Gompertz Model with Time-Dependent Covariate and Right Censored Data. *Sains Malaysiana*, 41(4)(2012), 471-480.
3. Kiani, K. and Arasan, J. (2012). Simulation of Interval Censored Data in Medical and Biological Studies. *International Journal of Modern Physics: Conference Series (IJMPCS), Proceedings of the International Conference on Mathematical and Computational Biology 2011*, 112-118, doi: 10.1142/S2010194512005168.
4. Kiani, K. and Arasan, J. (2012). Interval Estimations For Parameters Of Exponential Model With Doubly Interval-Censored Survival Time Data. In *Proceedings of the 2nd Regional Conference on Applied and Engineering Mathematics 2012 (RCAEM-II)*, 30-31 May, Penang, MALAYSIA.
5. Kiani, K., Arasan, J. and Midi, H. (2010). Confidence Interval Estimations for Parameters of Gompertz Survival Time Model. In *Proceedings of the Fundamental Science Congress: Mathematics Symposium (FSC 2010)*.
6. Kiani, K. and Arasan, J. (2009). Gompertz Model with Time-Dependent Covariate and Right Censored Data. In *Proceedings of the Symposium Kebangsaan Sains Matematik (KE-17)*, 15-17 December 2009, Melaka, MALAYSIA, pp: 771-777.